# EE/Ma 126b Information Theory - Homework Set \#6 

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6.1 The cooperative capacity of a multiple access channel.
(a) Since $X_{1}$ and $X_{2}$ know both indices, we can regard the pair of $\left(X_{1}, X_{2}\right)$ as one single codeword $X$. Thus for some distribution $p\left(x_{1}, x_{2}\right)$, or, distribution $p(x)$, we can design code $X(1), X(2), \ldots, X\left(2^{R_{1}+R_{2}}\right)$, i.i.d. $\sim p(x)$ and thus we will get the error $P_{e}^{(n)} \rightarrow 0$ when $n \rightarrow \infty$ if

$$
R_{1}+R_{2}<I\left(X_{1}, X_{2} ; Y\right) .
$$

(This is just the situation of one user channel.) And we also can get the converse that all code with error $P_{e}^{(n)} \rightarrow 0$ when $n \rightarrow \infty$ should have

$$
R_{1}+R_{2} \leq I\left(X_{1}, X_{2} ; Y\right)
$$

Thus we know that as taking $X_{1}$ and $X_{2}$ as a single codeword, the rate sum of $R_{1}$ and $R_{2}$ is bounded by

$$
\begin{equation*}
R_{1}+R_{2} \leq \max _{p\left(x_{1}, x_{2}\right)} I\left(X_{1}, X_{2} ; Y\right)=C \text {. } \tag{1}
\end{equation*}
$$

And since we can achieve $(C, 0)$ by setting $X_{1}=X$ and $X_{2}$ is null, or $(0, C)$ by setting $X_{1}$ null and $X_{2}=X$, we know the capacity region is just (1).
(b) For $Y=X_{1}+X_{2}, X_{i} \in\{0,1\}$, we get

$$
I\left(X_{1}, X_{2} ; Y\right)=H(Y)-H\left(Y \mid X_{1}, X_{2}\right)=H(Y) \leq \log 3
$$

since $Y \in\{0,1,2\}$. And we can achieve $\log 3$ by setting $p(0,0)=p(1,1)=\frac{1}{3}$ and $p(1,0)+p(0,1)=\frac{1}{3}$. Thus $C=\log 3$, and the capacity region is

$$
R_{1}+R_{2} \leq \log 3 \text {. }
$$

This region is larger than the one without cooperation between $X_{1}$ and $X_{2}$, since $\log 3>$ 1.5.
6.2 Capacity of multiple access channels.
(a) We can achieve the rate pair $\left(R_{1}, R_{2}\right)=(1,0)$ by setting $X_{2}=0$. Similarly, setting $X_{1}=0$, we can achieve rate pair $(0,1)$. For any distribution $p_{1}\left(x_{1}\right) p_{2}\left(x_{2}\right)$, we always have

$$
I\left(X_{1}, X_{2} ; Y\right) \leq H(Y) \leq 1,
$$

since $Y$ is binary. Thus the capacity region is just $R_{1}+R_{2} \leq 1$.
(b) If we use the 1-1 mapping $\{-1,1\} \rightarrow\{0,1\}:-1 \mapsto 1,1 \mapsto 0$ for both $X_{i}$ and $Y$, the function $Y=X_{1} \cdot X_{2}$ over $\{-1,1\}$ is exactly the function $Y=X_{1} \oplus X_{2}$ over $\{0,1\}$. Since for discrete channels, we do not care what the symbol is, the capacity region is exactly the same as that in (a). That is, $R_{1}+R_{2} \leq 1$.
6.3 Gaussian multiple access channel capacity. Let $f_{1}\left(x_{1}\right)$ be some normal distribution with variance $P_{1}-\epsilon$ and $f_{2}\left(x_{2}\right)$ be some normal distribution with variance $P_{2}-\epsilon$. The bold alphabets denote vectors of length $n$.

Codebook. Generate the codewords $\mathbf{X}_{1}(1), \mathbf{X}_{1}(2), \ldots, \mathbf{X}_{1}\left(2^{n R_{1}}\right)$ i.i.d. $\sim \prod_{i=1}^{n} f_{1}\left(x_{1, i}\right)$, and $\mathbf{X}_{2}(1), \mathbf{X}_{2}(2), \ldots, \mathbf{X}_{2}\left(2^{n R_{2}}\right)$ i.i.d. $\sim \prod_{i=1}^{n} f_{2}\left(x_{2, i}\right)$. These codewords form the codebook, which is revealed to the senders and the receiver.

Encoding. To send index $i$, sender 1 sends $\mathbf{X}_{1}(i)$. Similarly, to send $j$, sender 2 sends $\mathbf{X}_{2}(j)$.
Decoding. Let $A_{\epsilon}^{(n)}$ denote the set of typical $\left(\mathbf{X}_{1}, \mathbf{X}_{2}, \mathbf{Y}\right)$ sequences. Then the receiver chooses the pair $(i, j)$ such that

$$
\left(\mathbf{X}_{1}(i), \mathbf{X}_{2}(j), \mathbf{Y}\right) \in A_{\epsilon}^{(n)}
$$

if such a pair $(i, j)$ exists and is unique; otherwise, an error is declared.
Error probability. By symmetry, assume that pair $(1,1)$ is sent. Define the events

$$
E_{1}=\left\{\frac{1}{n} \sum_{i=1}^{n} X_{1, i}^{2}>P_{1}\right\}, E_{2}=\left\{\frac{1}{n} \sum_{i=1}^{n} X_{2, i}^{2}>P_{2}\right\},
$$

and

$$
E_{i, j}=\left\{\left(\mathbf{X}_{1}(i), \mathbf{X}_{2}(j), \mathbf{Y}\right) \in A_{\epsilon}^{(n)}\right\} .
$$

Then an error occurs if $E_{1}$ or $E_{2}$ occurs (the power constraints are violated) or $E_{1,1}^{c}$ occurs (the transmitted codewords and the received sequence are not jointly typical) or $E_{i, j}$ with $(i, j) \neq(1,1)$ occurs (some wrong codewords is jointly typical with the received sequence). Let $P$ denote the expectation over all codebooks of the conditional probability given $(1,1)$ sent. That is, for some event $\mathcal{E}$,

$$
P(\mathcal{E})=E \operatorname{Pr}\{\mathcal{E} \mid(1,1) \text { sent }\}
$$

Thus

$$
\begin{aligned}
E P_{e}^{(n)} & =P\left(E_{1} \cup E_{2} \cup E_{1,1}^{c} \cup \bigcup_{(i, j) \neq(1,1)} E_{i, j}\right) \\
& \leq P\left(E_{1}\right)+P\left(E_{2}\right)+P\left(E_{1,1}^{c}\right)+\sum_{i \neq 1} P\left(E_{i, 1}\right)+\sum_{j \neq 1} P\left(E_{1, j}\right)+\sum_{i \neq 1} \sum_{j \neq 1} P\left(E_{i, j}\right) .
\end{aligned}
$$

By the WLLN, $P\left(E_{1}\right) \rightarrow 0$ and $P\left(E_{2}\right) \rightarrow 0$ as $n \rightarrow \infty$. By the joint AEP, $P\left(E_{1,1}^{c}\right) \leq$ $\epsilon$ for $n$ sufficiently large. Since by the code generation process, $\mathbf{X}_{1}(1)$ and $\mathbf{X}_{1}(i)$ are independent $(i \neq 1)$, so are $\mathbf{Y}$ and $\mathbf{X}_{1}(i)$ given $\mathbf{X}_{2}(1)$. Hence by the joint AEP, $P\left(E_{i, 1}\right) \leq$
$2^{-n\left(I\left(X_{1} ; Y \mid X_{2}\right)-\epsilon\right)}$. Similarly, $P\left(E_{1, j}\right) \leq 2^{-n\left(I\left(X_{2} ; Y \mid X_{1}\right)-\epsilon\right)}$. And for $i \neq 1$ and $j \neq 1$, $\mathbf{X}_{1}(i), \mathbf{X}_{2}(j)$ and $\mathbf{Y}$ are independent, so by joint AEP $P\left(E_{i, j}\right) \leq 2^{-n\left(I\left(X_{1}, X_{2} ; Y\right)-\epsilon\right)}$. Thus

$$
\begin{aligned}
E P_{e}^{(n)} \leq & 3 \epsilon+2^{-n\left(I\left(X_{1} ; Y \mid X_{2}\right)-R_{1}-\epsilon\right)}+2^{-n\left(I\left(X_{2} ; Y \mid X_{1}\right)-R_{2}-\epsilon\right)} \\
& +2^{-n\left(I\left(X_{1}, X_{2} ; Y\right)-R_{1}-R_{2}-\epsilon\right)} \\
\leq & 6 \epsilon
\end{aligned}
$$

for $n$ sufficiently large and

$$
\begin{align*}
R_{1} & <I\left(X_{1} ; Y \mid X_{2}\right)  \tag{2}\\
R_{2} & <I\left(X_{2} ; Y \mid X_{1}\right)  \tag{3}\\
R_{1}+R_{2} & <I\left(X_{1}, X_{2} ; Y\right) . \tag{4}
\end{align*}
$$

This proves the existence of good $\left(\left(2^{n R_{1}}, 2^{n R_{2}}\right), n\right)$ code. And by choosing a good codebook and deleting the worst half of both sets of codewords, we obtain a code with low maximal probability of error. In particular, the power constraints are satisfied by each of the remaining codewords (similar to to proof for the single user Gaussian channel).

The bounds (2), (3) and (4) give the achievable region for a specific distribution. The full achievable region is the closure of the convex hull of those regions for any distribution. Since $X_{1}$ and $X_{2}$ are independent (by the code generation), and the noise $Z$ is also independent of $X_{1}$ and $X_{2}$, we have

$$
\begin{equation*}
I\left(X_{1} ; Y \mid X_{2}\right)=I\left(X_{1} ; X_{1}+X_{2}+Z \mid X_{2}\right)=h\left(X_{1}+Z\right)-h\left(X_{1}\right) \leq \frac{1}{2} \log \left(1+\frac{P_{1}}{N}\right) \tag{5}
\end{equation*}
$$

and we can achieve the equality by setting $X_{1} \sim \mathcal{N}\left(0, P_{1}\right)$. Similarly, $I\left(X_{2} ; Y \mid X_{1}\right) \leq \frac{1}{2} \log \left(1+\frac{P_{2}}{N}\right)$ with equality when $X_{2} \sim \mathcal{N}\left(0, P_{2}\right)$. And

$$
\begin{equation*}
I\left(X_{1}, X_{2} ; Y\right)=I\left(X_{1}, X_{2} ; X_{1}+X_{2}+Z\right)=h\left(X_{1}+X_{2}+Z\right)-h(Z) \leq \frac{1}{2} \log \left(1+\frac{P_{1}+P_{2}}{N}\right) \tag{6}
\end{equation*}
$$

with equality when $X_{1} \sim \mathcal{N}\left(0, P_{1}\right)$ and $X_{2} \sim \mathcal{N}\left(0, P_{2}\right)$. So finally, we get the achievable region is

$$
\begin{aligned}
R_{1} & <\frac{1}{2} \log \left(1+\frac{P_{1}}{N}\right) \\
R_{2} & <\frac{1}{2} \log \left(1+\frac{P_{2}}{N}\right) \\
R_{1}+R_{2} & <\frac{1}{2} \log \left(1+\frac{P_{1}+P_{2}}{N}\right) .
\end{aligned}
$$

6.4 Converse for the Gaussian multiple access channel. Consider any $\left(\left(2^{n R_{1}}, 2^{n R_{2}}\right), n\right)$ code that
satisfies the power constraints and with error probability $P_{e}^{(n)} \rightarrow 0$ when $n \rightarrow \infty$. Then

$$
\begin{aligned}
n R_{1} & =H\left(W_{1}\right)=I\left(W_{1} ; \mathbf{Y}\right)+H\left(W_{1} \mid \mathbf{Y}\right) \\
& \leq I\left(W_{1} ; \mathbf{Y}\right)+n \epsilon_{n} \\
& \leq I\left(\mathbf{X}_{1}\left(W_{1}\right) ; \mathbf{Y}\right)+n \epsilon_{n} \\
& =h\left(\mathbf{X}_{1}\left(W_{1}\right)\right)-h\left(\mathbf{X}_{1}\left(W_{1}\right) \mid \mathbf{Y}\right)+n \epsilon_{n} \\
& \leq h\left(\mathbf{X}_{1}\left(W_{1}\right) \mid \mathbf{X}_{2}\left(W_{2}\right)\right)-h\left(\mathbf{X}_{1}\left(W_{1}\right) \mid \mathbf{Y}, \mathbf{X}_{2}\left(W_{2}\right)\right)+n \epsilon_{n} \\
& =I\left(\mathbf{X}_{1}\left(W_{1}\right) ; \mathbf{Y} \mid \mathbf{X}_{2}\left(W_{2}\right)\right)+n \epsilon_{n} \\
& =h\left(\mathbf{Y} \mid \mathbf{X}_{2}\left(W_{2}\right)\right)-h\left(\mathbf{Y} \mid \mathbf{X}_{1}\left(W_{1}\right), \mathbf{X}_{2}\left(W_{2}\right)\right)+n \epsilon_{n} \\
& =h\left(\mathbf{Y} \mid \mathbf{X}_{2}\left(W_{2}\right)\right)-\sum_{i=1}^{n} h\left(Y_{i} \mid Y_{1}^{i-1}, \mathbf{X}_{1}\left(W_{1}\right), \mathbf{X}_{2}\left(W_{2}\right)\right)+n \epsilon_{n} \\
& =h\left(\mathbf{Y} \mid \mathbf{X}_{2}\left(W_{2}\right)\right)-\sum_{i=1}^{n} h\left(Y_{i} \mid X_{1, i}, X_{2, i}\right)+n \epsilon_{n} \\
& \leq \sum_{i=1}^{n} h\left(Y_{i} \mid \mathbf{X}_{2}\left(W_{2}\right)\right)-\sum_{i=1}^{n} h\left(Y_{i} \mid X_{1, i}, X_{2, i}\right)+n \epsilon_{n} \\
& \leq \sum_{i=1}^{n} h\left(Y_{i} \mid X_{2, i}\right)-\sum_{i=1}^{n} h\left(Y_{i} \mid X_{1, i}, X_{2, i}\right)+n \epsilon_{n} \\
& =\sum_{i=1}^{n} I\left(X_{1, i} ; Y_{i} \mid X_{2, i}\right)+n \epsilon_{n} .
\end{aligned}
$$

This is very similar to the proof in the discrete case and all the reasons are the same (though in continuous version). Now let $P_{1, i}$ be the average power of the $i$ th codeword $X_{1, i}$, i.e.,

$$
P_{1, i}=\frac{1}{2^{n R_{1}}} \sum_{w} X_{1, i}^{2}(w) .
$$

Then since $Y_{i}=X_{1, i}+X_{2, i}+Z_{i}$, and $X_{1, i}, X_{2, i}$ and $Z_{i}$ are independent (since $W_{1}$ and $W_{2}$ are independent and $Z$ are independent of $\left(X_{1}, X_{2}\right)$ ), we have (similar to (5))

$$
I\left(X_{1, i} ; Y_{i} \mid X_{2, i}\right)=h\left(X_{1, i}+Z_{i}\right)-h\left(X_{1, i}\right) \leq \frac{1}{2} \log \left(1+\frac{P_{1, i}}{N}\right) .
$$

Continuing with the inequalities of the converse, we get

$$
\begin{aligned}
R_{1} & \leq \frac{1}{n} \sum_{i=1}^{n} \frac{1}{2} \log \left(1+\frac{P_{1, i}}{N}\right)+\epsilon_{n} \\
& \leq \frac{1}{2} \log \left(1+\frac{1}{n} \sum_{i=1}^{n} \frac{P_{1, i}}{N}\right)+\epsilon_{n} \\
& \leq \frac{1}{2} \log \left(1+\frac{P_{1}}{N}\right)+\epsilon_{n},
\end{aligned}
$$

since $\frac{1}{2} \log (1+x)$ is a concave function of $x$ and we also have the power constraint $\frac{1}{n} \sum_{i} P_{1, i} \leq P_{1}$. Together with $\epsilon_{n} \rightarrow 0$ when $n \rightarrow \infty$, we have

$$
R_{1} \leq \frac{1}{2} \log \left(1+\frac{P_{1}}{N}\right) .
$$

Similarly, we have

$$
R_{2} \leq \frac{1}{2} \log \left(1+\frac{P_{2}}{N}\right)
$$

To bound the sum of the rates, we have (oh, I won't write all the details. They are just an analog to the proof in the discrete case.)

$$
n\left(R_{1}+R_{2}\right) \leq \sum_{i=1}^{n} I\left(X_{1, i}, X_{2, i} ; Y_{i}\right)+n \epsilon_{n}
$$

With (6), we obtain $I\left(X_{1, i}, X_{2, i} ; Y_{i}\right) \leq \frac{1}{2} \log \left(1+\frac{P_{1, i}+P_{2, i}}{N}\right)$. Thus

$$
\begin{aligned}
R_{1}+R_{2} & \leq \frac{1}{n} \sum_{i=1}^{n} \frac{1}{2} \log \left(1+\frac{P_{1, i}+P_{2, i}}{N}\right)+\epsilon_{n} \\
& \leq \frac{1}{2} \log \left(1+\frac{1}{n} \sum_{i=1}^{n} \frac{P_{1, i}+P_{2, i}}{N}\right)+\epsilon_{n} \\
& \leq \frac{1}{2} \log \left(1+\frac{P_{1}+P_{2}}{N}\right)+\epsilon_{n}
\end{aligned}
$$

Since $\epsilon_{n} \rightarrow 0$ when $n \rightarrow \infty$, we get

$$
R_{1}+R_{2} \leq \frac{1}{2} \log \left(1+\frac{P_{1}+P_{2}}{N}\right)
$$

This finally finishes the proof of the converse.

### 6.5 Unusual multiple access channel.

(a) If we set $X_{1}=0$, the channel becomes $Y=X_{2}$ and we can send $X_{2}$ at a rate $R_{2}=1$. Thus the rate pair $(0,1)$ is achievable. Symmetrically, the pair $(1,0)$ is also achievable.
(b) For any non-degenerate distribution $p_{1}\left(x_{1}\right) p_{2}\left(x_{2}\right)$, we have

$$
\begin{aligned}
I\left(X_{1}, X_{2} ; Y\right) & =H(Y)-H\left(Y \mid X_{1}, X_{2}\right) \\
& =H(Y)-\sum_{x_{1}, x_{2}} H\left(Y \mid X_{1}=x_{1}, X_{2}=x_{2}\right) p_{1}\left(x_{1}\right) p_{2}\left(x_{2}\right) \\
& =H(Y)-H\left(Y \mid X_{1}=1, X_{2}=1\right) p_{1}(1) p_{2}(1) \\
& =H(Y)-p_{1}(1) p_{2}(1) \\
& <1
\end{aligned}
$$

The reason for the inequality is that $Y$ is binary so $H(Y) \leq 1$, and $p_{1}(1) p_{2}(1)>0$.
(c) Consider the rate pair $\left(\frac{1}{2}, \frac{1}{2}\right)$. It can be achieved by timesharing of those two rate pairs $(1,0)$ and $(0,1)$. However, for any non-degenerate distribution $p_{1}\left(x_{1}\right) p_{2}\left(x_{2}\right)$, we have $I\left(X_{1}, X_{2} ; Y\right)<1$, and for any degenerate distribution, we have $I\left(X_{1} ; Y \mid X_{2}\right)=0$ and/or $I\left(X_{2} ; Y \mid X_{1}\right)=0$. Thus the pair $\left(\frac{1}{2}, \frac{1}{2}\right)$ can not be within the region defined by

$$
\begin{aligned}
R_{1} & \leq I\left(X_{1} ; Y \mid X_{2}\right) \\
R_{2} & \leq I\left(X_{2} ; Y \mid X_{1}\right) \\
R_{1}+R_{2} & \leq I\left(X_{1}, X_{2} ; Y\right)
\end{aligned}
$$

for any $p_{1}\left(x_{1}\right) p_{2}\left(x_{2}\right)$.

